

Large Deviations and Weak Convergence of Measures

with applications to Monte Carlo Estimators

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Motivation

- Consider a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and real valued random variable $X : \Omega \rightarrow \mathbb{R}$, with distribution $\mu := \mathbb{P} \circ X^{-1}$.



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- In many applications we are interested in computing the expected value $\mathbb{E}[X]$

$$\theta = \mathbb{E}[X] = \int_{\Omega} X \, d\mathbb{P}.$$

- In practice, it may not be possible to compute this integral and Monte Carlo (MC) methods are often used to simulate θ .



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- *Main idea behind MC-method:* Strong law of large numbers (SLLN) implies that $\theta_n(\omega) \rightarrow \theta$ almost surely.



Convergence rate of the CMC estimator

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- If we want $\mathbb{P}(|\theta_n - \theta| < \varepsilon|\theta|)$, then we need

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- For rare events the variance and expectation are almost the same!

$$\frac{\mathbb{V}[\mathbf{1}_A]}{\mathbb{E}[\mathbf{1}_A]} = 1 - p, \quad n \gg \frac{z_{1-\alpha/2}^2}{\varepsilon^2 p}$$



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- First mathematical results in the theory of large deviations was published in 1938 [1], by Harald Cramér (actuary and affiliated with Stockholm University).
- Cramér's Motivation was insurance mathematics and ruin probabilities.
- S.R.S. Varadhan introduced the modern mathematical theory of large deviations. Seminal paper: [5] (Abel Prize for his contributions).



On a new limit theorem in probability theory
(Sur un nouveau théorème-limite de la théorie des probabilités)

Harald Cramér (1893–1985)
Stockholm, Sweden

Translated by
Hugo Touchette
National Institute for Theoretical Physics (NITheP), Stellenbosch, South Africa
15 March 2018

Original article: H. Cramér, Sur un nouveau théorème-limite de la théorie des probabilités, Colloque consacré à la théorie des probabilités, Actualités scientifiques et industrielles 716, 2–25, Hermann & Co, Paris, 1938.

Reprinted in: H. Cramér, *Collected Works*, A. Martin-Löf (Ed.), Vol. II, Springer, Berlin, 1994, p. 395–413.



Chapitre premier

Considérons une suite Z_1, Z_2, \dots de variables aléatoires indépendantes ayant toutes la même fonction de répartition $F(x)$, et telle que

$$F(Z) = 0, \quad F(Z_0^2) = \sigma^2 > 0. \quad (1)$$

Désignons par $W_n(x)$ la fonction de répartition de la somme

$$Z_1 + \dots + Z_n,$$

et par $F_n(x)$ la fonction de répartition de la variable

$$\frac{Z_1 + \dots + Z_n}{\sigma \sqrt{n}}.$$

On a donc

$$F_n(x) = \text{Prob}(Z_1 + \dots + Z_n \leq \sigma \sqrt{n} x)$$

et

$$F_n(x) = W_n(\sigma \sqrt{n} x). \quad (2)$$

D'après la théorie limite classique de Laplace-Legendre (sans la forme moderne précisée par Lindberg et par M. Paul Lévy) on a alors pour chaque valeur réelle fixe de x

$$\lim_{n \rightarrow \infty} F_n(x) = \Phi(x) = \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{1}{2}t^2} dt. \quad (3)$$

Pour ce théorème, on a donc une expression asymptotique (pour $n \rightarrow \infty$) de la probabilité $F_n(x)$ de l'angle $Z_1 + \dots + Z_n \leq \sigma \sqrt{n} x$

$$Z_1 + \dots + Z_n \leq \sigma \sqrt{n} x$$

ou, ce qui revient au même, de la probabilité $1 - F_n(x)$ de l'angle

$$Z_1 + \dots + Z_n > \sigma \sqrt{n} x$$

à deux termes un nombre réel indépendant de n .

Il est alors naturel de se demander ce que deviennent ces probabilités lorsque n peut varier avec n , en restant vers $+\infty$ ou $-\infty$, quand n croît indéfiniment.

Dans ces conditions, la fonction (1) donne que le résultat équivaut

$$\lim_{n \rightarrow \infty} F_n(x) = \begin{cases} 1 & \text{quand } x \rightarrow +\infty, \\ 0 & \text{quand } x \rightarrow -\infty, \end{cases}$$

qui exprime seulement que $F_n(x)$ tend vers les mêmes limites que $\Phi(x)$ lorsque $n \rightarrow \infty$.

Pour savoir si l'équivalence asymptotique de $F_n(x)$ et $\Phi(x)$ subsiste dans les conditions indiquées, on peut se proposer d'étudier les rapports

$$\frac{1 - F_n(x)}{1 - \Phi(x)} \quad \text{pour } x \rightarrow +\infty, \quad (4)$$

First chapter

Consider a sequence Z_1, Z_2, \dots of independent random variables having the same cumulative distribution function $F(x)$ and such that

$$F(Z) = 0, \quad F(Z_0^2) = \sigma^2 > 0. \quad (1)$$

Denote by $W_n(x)$ the cumulative distribution function of the sum

$$Z_1 + \dots + Z_n,$$

and by $F_n(x)$ the cumulative distribution function of the variable

$$\frac{Z_1 + \dots + Z_n}{\sigma \sqrt{n}}.$$

We [2] thus have

$$F_n(x) = \text{Prob}(Z_1 + \dots + Z_n \leq \sigma \sqrt{n} x)$$

and

$$F_n(x) = W_n(\sigma \sqrt{n} x). \quad (2)$$

Following the classical limit theorem of Laplace-Legendre [3] (in its modern version specified by Lindberg and by Paul Lévy) we thus have for each real value x

$$\lim_{n \rightarrow \infty} F_n(x) = \Phi(x) = \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{1}{2}t^2} dt. \quad (3)$$

From this theorem [4], we thus have an asymptotic expression (for $n \rightarrow \infty$) for the probability $F_n(x)$ of the inequality

$$Z_1 + \dots + Z_n \leq \sigma \sqrt{n} x$$

or, which amounts to the same, for the probability $1 - F_n(x)$ of the inequality

$$Z_1 + \dots + Z_n > \sigma \sqrt{n} x$$

to being as before a real number independent of n .

It is thus natural to ask what happens of these probabilities when n can vary with n , going to $+\infty$ or $-\infty$ when n grows indefinitely.

In these conditions, Relation (1) only gives the evident result

$$\lim_{n \rightarrow \infty} F_n(x) = \begin{cases} 1 & \text{when } x \rightarrow +\infty, \\ 0 & \text{when } x \rightarrow -\infty, \end{cases}$$

which expresses only that $F_n(x)$ converges to the same limits as $\Phi(x)$ when $n \rightarrow \infty$.

To see whether the asymptotic equivalence of $F_n(x)$ and $\Phi(x)$ remains under the mentioned conditions indicated, we could propose to study the ratios

$$\frac{1 - F_n(x)}{1 - \Phi(x)} \quad \text{when } x \rightarrow +\infty, \quad (4)$$

Figure: Translation of Cramér's publication from French to English by Hugo Touchette [2]



Large Deviations

- A sequence (X_n) of i.i.d. random variables taking values in a Hausdorff topological space \mathcal{X} .
- In the theory of large deviations we want to find a rate function $I : \mathcal{X} \rightarrow [0, \infty]$, such that

$$\lim_{n \rightarrow \infty} \mathbb{P}(X_n \in A) \approx e^{-n \inf_{x \in A} I(x)}.$$

- Can be used to approximate rare event probabilities.



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- Can be used to approximate rare event probabilities.
- Can be used to analyze the convergence of MC estimators (substitute $X_n = \theta_n$).



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- Let $X : \Omega \rightarrow \mathcal{X}$ be a random variable. If $f : \mathcal{X} \rightarrow \mathbb{R}$ is measurable, then $f(X)$ is a real valued random variable.



Empirical Distributions

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$$\delta : \Omega \rightarrow \mathbf{M}_1(\mathcal{X}), \quad \delta_{X(\omega)}(A) := \begin{cases} 1, & X(\omega) \in A, \\ 0, & X(\omega) \notin A, \end{cases}$$



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- Integrating over \mathcal{X} with respect to the measure $\mathbf{L}_n(\omega)$:

$$\int_{\mathcal{X}} f \, d\mathbf{L}_n(\omega) = \frac{1}{n} \sum_{i=1}^n \int_{\mathcal{X}} f \, d\delta_{X_i(\omega)} = \frac{1}{n} \sum_{i=1}^n f(X_i(\omega))$$



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- Does \mathbf{L}_n converge to the distribution μ of \mathcal{X} in $\mathbf{M}_1(\mathcal{X})$?



- Convergence is a topological concept.



Topologies on $\mathbf{M}(\mathcal{X})$

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Three important spaces of measures:

- 1 $\mathbf{M}(\mathcal{X})$ finite signed measures on \mathcal{X} (is a linear space)
- 2 $\mathbf{M}_+(\mathcal{X})$ nonnegative finite measures on \mathcal{X}
- 3 $\mathbf{M}_1(\mathcal{X})$ probability measures on \mathcal{X}



The τ -topology

- Let $f : \mathcal{X} \rightarrow \mathbb{R}$ be a bounded measurable function and $\mu \in \mathbf{M}(\mathcal{X})$.
Then

$$\langle f, \mu \rangle = \int_{\mathcal{X}} f \, d\mu,$$

is a dual pairing.



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- The maps $\langle f, \cdot \rangle : \mathbf{M}(\mathcal{X}) \rightarrow \mathbb{R}$ generate a weak topology on $\mathbf{M}(\mathcal{X})$:
the τ -topology
- μ_α converges to μ in the τ -topology iff

$$\lim_{\alpha} \int_{\mathcal{X}} f \, d\mu_{\alpha} = \int_{\mathcal{X}} f \, d\mu,$$

for every bounded measurable function f .



The topology of weak convergence

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Definition

Let \mathcal{X} be a metrizable space, then a net (μ_α) in $\mathbf{M}(\mathcal{X})$ converges weakly to $\mu \in \mathbf{M}(\mathcal{X})$ if

$$\lim_{\alpha} \int_{\mathcal{X}} f d\mu_{\alpha} = \int_{\mathcal{X}} f d\mu, \quad \text{for every } f \in C_b(\mathcal{X}).$$



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The topology of weak convergence is weaker (has less open sets) than the τ -topology!



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- Extend many results from $\mathbf{M}_1(\mathcal{X})$ to $\mathbf{M}_+(\mathcal{X})$.



- Weak convergence of measures is often used in probability theory.
- In Ch. 3 of the thesis we study the the τ -topology and the topology of weak convergence.
- Extend many results from $\mathbf{M}_1(\mathcal{X})$ to $\mathbf{M}_+(\mathcal{X})$.
- Builds upon the work of Varadarajan in [4].



Convergence of MC-estimators

Varadarajan proved in 1958 the empirical distributions of the CMC-estimator converges weakly [3].



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Theorem

Let \mathcal{X} be a separable metrizable space and (X_i) a sequence of i.i.d. random variables taking values in \mathcal{X} with law μ . Then the empirical distributions \mathbf{L}_n converge weakly to μ almost surely, i.e.

$$\mathbb{P}(\{\omega \in \Omega : \mathbf{L}_n(\omega) \Longrightarrow \mu\}) = 1.$$

We extend this result to the empirical distribution of the importance sampling estimator.



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$$\{x \in \mathcal{X} : f(x) \leq t\}, \quad t \in [0, \infty).$$



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A sequence of probability measures (μ_n) the *large deviation principle (LDP)* with rate function I if

$$-\inf_U I \leq \liminf_{n \rightarrow \infty} \frac{1}{n} \log [\mu_n(U)]$$

for every open set U , and

$$-\inf_C I \geq \limsup_{n \rightarrow \infty} \frac{1}{n} \log [\mu_n(C)]$$

for every closed set C .



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- 3 Identification of the rate function.



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- **The LDP is unique!**
- The upper bound can be hard to prove...
- (μ_n) satisfies a *weak large deviation principle* with rate function I if it satisfies the lower bound and

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for every compact set $K \subset \mathcal{X}$.



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for every compact set $K \subset \mathcal{X}$.

- Easier to prove a weak LDP. Can go from weak to full LDP by proving that the sequence (μ_n) is exponentially tight.



Large Deviations of empirical means

- Consider (X_i) i.i.d. random variables with distribution μ defined on $(\Omega, \mathcal{F}, \mathbb{P})$ and taking values in the topological linear space \mathcal{X} .



Large Deviations of empirical means

- Consider (X_i) i.i.d. random variables with distribution μ defined on $(\Omega, \mathcal{F}, \mathbb{P})$ and taking values in the topological linear space \mathcal{X} .
- The empirical means $\mathbf{S}_n : \Omega \rightarrow \mathcal{X}$, defined by

$$\mathbf{S}_n(\omega) = \frac{1}{n} \sum_{i=1}^n X_i(\omega).$$

- The distributions are given by $\mu_n = \mathbb{P} \circ \mathbf{S}_n^{-1}$.



Theorem (Weak Cramér's Theorem)

The sequence (μ_n) of distributions of the empirical means satisfy a weak large deviation principle with a convex rate function $I = \Lambda^*$, and

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \mu_n(A) = - \inf_{x \in A} \Lambda^*(x),$$

for every convex and open $A \subset \mathcal{X}$.

Here Λ^* is the Legendre-Fenchel transform of

$$\Lambda(\lambda) = \Lambda_\mu := (\lambda) = \log \mathbb{E} \left[e^{\langle \lambda, X \rangle} \right] = \log \left[\int_{\mathcal{X}} e^{\langle \lambda, d \rangle} d\mu(x) \right].$$



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$$\mathbb{P}\left(|\theta_n - \theta| < \varepsilon|\theta|\right) > 1 - \alpha.$$



Large Deviations and CMC

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- Let

$$R_\varepsilon := B(\theta, \varepsilon|\theta|), \quad A_\varepsilon = R_\varepsilon^c,$$



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- Let

$$R_\varepsilon := B(\theta, \varepsilon|\theta|), \quad A_\varepsilon = R_\varepsilon^c,$$

then we get

$$\mathbb{P}(\theta_n \in A_\varepsilon) \leq \alpha.$$



Large Deviations and CMC

- If we want

$$\mathbb{P}\left(|\theta_n - \theta| < \varepsilon|\theta|\right) > 1 - \alpha.$$

- Let

$$R_\varepsilon := B(\theta, \varepsilon|\theta|), \quad A_\varepsilon = R_\varepsilon^c,$$

then we get

$$\mathbb{P}(\theta_n \in A_\varepsilon) \leq \alpha.$$

- For large n we can interpret this as

$$\mu_n(A_\varepsilon) \lesssim e^{-nI(A_\varepsilon)}.$$



- To achieve the desired precision with confidence $1 - \alpha$ we want

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- Rearranging yields:

$$n \gtrsim \frac{\log(\alpha)}{I(A_\varepsilon)}.$$



Example with $X \sim N(\theta, \sigma^2)$

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- Compare with introduction:

$$n \gtrsim \frac{z_{1-\alpha/2}^2\sigma^2}{\varepsilon^2\theta^2}.$$



Definition

Let $f : \mathcal{X} \rightarrow [-\infty, \infty]$, then the *Legendre-Fenchel transform* of f is the function $f^* : \mathcal{X}^* \rightarrow [-\infty, \infty]$ defined by

$$\begin{aligned} f^*(\lambda) &= \sup\{\langle \lambda, x \rangle - f(x) : x \in \mathcal{X}\} \\ &= -\inf\{f(x) - \langle \lambda, x \rangle : x \in \mathcal{X}\}. \end{aligned}$$



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Theorem (Biconjugate Theorem)

Let $f : \mathcal{X} \rightarrow (-\infty, \infty]$ not be identically ∞ , then $f = f^{**}$ if and only if f is convex and lower semicontinuous.



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Questions

Feel free to ask any questions!



- [1] Harald Cramér. “Sur un nouveau theoreme-limite de la theorie des probabilities”. In: *Scientifiques et Industrielles* 736 (1938), pp. 5–23.
- [2] Harald Cramér and Hugo Touchette. *On a new limit theorem in probability theory (Translation of 'Sur un nouveau théorème-limite de la théorie des probabilités')*. 2022. arXiv: 1802.05988 [math.HO].
- [3] Veeravalli S Varadarajan. “On the convergence of sample probability distributions”. In: *Sankhyā: The Indian Journal of Statistics (1933-1960)* 19.1/2 (1958), pp. 23–26.
- [4] Veeravalli S Varadarajan. “Weak convergence of measures on separable metric spaces”. In: *Sankhyā: The Indian Journal of Statistics (1933-1960)* 19.1/2 (1958), pp. 15–22.



- [5] S. R. S. Varadhan. “Asymptotic probabilities and differential equations”. In: *Communications on Pure and Applied Mathematics* 19.3 (1966), pp. 261–286. DOI: <https://doi.org/10.1002/cpa.3160190303>. eprint: <https://onlinelibrary.wiley.com/doi/pdf/10.1002/cpa.3160190303>. URL: <https://onlinelibrary.wiley.com/doi/abs/10.1002/cpa.3160190303>.

